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Impacts of Trade and the Environment on Clustered Multilateral Environmental Agreements

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1. INTRODUCTION

The interconnectedness of the global environment is beyond dispute. [...] coordinated international action is essential to protecting Earth's climate, preserving its biodiversity, and managing its marine and other common resources. (World Resources Institute, International Environmental Governance, 2003, p. 137)

PENNESS to the international economy and the environment intersects in many different Ways. While one part of the literature argues that a more stringent environmental policy harms goods trade (see, among others, d'Arge, 1974; Pethic, 1976; Siebert 1977; McGuire, 1982; OECD, 1985; Markusen, 1999; Ulph, 1999; Millimet and List, 2004) and foreign direct investment (see Walter, 1982; Pearson, 1985, 1987; Leonard, 1988; Taylor, 2005), the other part of the literature provides evidence of a limited detrimental impact of environmental policy for trade (see Tobey, 1990; Porter, 1991; van Beers and van den Bergh, 1997, 2000; Berman and Bui, 2001; Alpay et al., 2002; Lanoie et al., 2008) or foreign direct investment (see Porter and van der Linde, 1995; Bjørn et al., 1997; Xu and Song, 2000). Clearly, in the absence of evidence for strong negative economic effects of a more stringent environmental policy, one would even less hesitantly argue in favour of its inception. Multilateral environmental agreements (MEAs) are one widely adopted measure to introduce and coordinate stringent environmental policies across countries. Recent work at the interface between macroeconomics and international economics indicated that cooperation in environmental agreements may induce positive effects on international economic issues beyond the environment (see Rose and Spiegel, 2009) and cooperation in trade and investment agreements makes cooperation in environmental agreements more likely (see Egger et al., 2011).

While earlier academic work provided insights into the dynamic, economic and political determinants of participation in environmental agreements at large from a theoretical as well as an empirical perspective (see Chandler and Tulkens, 1992; Hoel, 1992, 2005; Barrett, 1994; Hoel and Schneider, 1997; Tol, 1997, 2001; Carraro, 1998; Finus and Rundshagen, 1998a, 1998b; Neumayer, 2002; Barrett and Stavins, 2003; Beron et al., 2003; Carraro et al., 2003; Eyckmans and Tulkens, 2003; Lange and Vogt, 2003; Mitchell, 2003; Murdoch et al., 2003; Buchholz et al., 2005; Carraro et al., 2006; Weikard et al., 2006; Egger et al., 2011), little is known about the role

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of these determinants for MEAs of specific type and about spillovers in the MEA participation across issues. Even at the level of all MEAs, Mitchell (2003, p. 431) emphasises that 'the empirical basis for claims regarding the number of such agreements and their characteristics remains weak'. Close to nothing is known about the determinants of the numbers of such agreements at the level of different environmental issues.

This paper aims at filling this gap by proposing an empirical model of the number of specific MEAs regarding environmental issues, which involves economic, political and environmental determinants. For this purpose, we classify MEAs into clusters covering different environmental issues. Following the classification of MEAs of the United Nations Environment Programme (2001; UNEP MEA Clusters), we discern five clusters: biodiversity, atmosphere, land, chemicals and hazardous wastes, and seas. Besides economic and political fundamentals of MEA ratification by cluster, we consider spillover effects in the MEA participation in some clusters on the participation in other clusters. This analysis reveals that MEA participation is not contagious across all clusters. For example, land-related MEAs only influence participation in biodiversity-specific MEAs, but not other ones. By way of contrast, MEAs classified as chemicals and hazardous wastes provide significant stimuli on the participation in MEAs on other issues.

The remainder of the paper is organised as follows. Section 2 describes MEA ratification across clusters and countries by way of maps. This provides first insights concerning the differences and similarities between clusters. Section 3 outlines the econometric model. We distinguish three key types of determinants of MEAs in Section 4, define the corresponding variables and report on the data sources. Section 5 summarises the parameter estimates and marginal effects based on dynamic linear feedback models for counts of MEAs in a cluster countries participate in. The last section offers some concluding remarks.

2. THE GEOGRAPHY OF MEA PARTICIPATION ACROSS CLUSTERS

We illustrate the geography of MEA ratifications across 199 countries and MEA clusters in the year 2006 by means of maps. For this, we compile data on MEAs from two sources: the Socioeconomic Data and Applications Center (SEDAC) Database of the Center for International Earth Science Information Network (CIESIN) (2006) and data provided by courtesy of Mitchell (2007). The latter database enabled to broaden the CIESIN SEDAC data set and extend it until the year 2006. For illustration, we associate specific grey shades with the 0–20, 20–40, 40–60, 60–80 and 80–100 quantiles of MEA numbers, with darker shades identifying relatively fewer and lighter shades identifying relatively more MEAs concluded in a country and cluster. The darkest two shades of grey indicate countries with fewer MEAs concluded in a cluster than the centre quintile of the distribution. Ordinary grey identifies countries with as many MEAs as in the centre quintile of the distribution. The lighter two shades of grey identify countries more MEAs ratified than the median country. Subsequently, we provide individual maps of MEAs in 2006 for the clusters biodiversity, atmosphere, land, chemicals and hazardous wastes, and seas.

From the colour scales and associated MEA numbers in the lower left corner of each of the five maps, it is evident that the MEA clusters differ with respect to the number of MEAs across the five quintiles. Only up to 21 MEAs in the cluster land were ratified by an economy, while up to 94 MEAs were signed by a country in the cluster seas. In the cluster Biodiversity, both France and Sweden ratified the largest number of MEAs until 2006, whereas Germany and Luxembourg were the ones with the highest number of MEAs in the atmosphere cluster. Germany also ratified the largest number of MEAs in the clusters *Land* and chemicals and hazardous wastes. France participated in more MEAs of the seas type than

any other economy in 2006. Besides European countries, the US, Canada, Japan, Korea, India, Australia, New Zealand, Russia, Egypt, Argentina, Chile, Ecuador and Peru were among the countries that had an above-average number of MEAs ratified (light grey or very light grey) across all types of MEA clusters (Figures 1–5).

3. ECONOMETRIC MODEL

While Figures 1–5 portrayed the number of MEAs per cluster and country as of 2006, it is important to note that the time process of a country's MEA ratification of any type displays



FIGURE 1 The Number of Multilateral Environmental Agreements in 2006 – Cluster: Biodiversity

FIGURE 2 The Number of Multilateral Environmental Agreements in 2006 – Cluster: Atmosphere





FIGURE 3 The Number of Multilateral Environmental Agreements in 2006 – Cluster: Land

FIGURE 4 The Number of Multilateral Environmental Agreements in 2006 – Cluster: Chemicals and Hazardous Wastes



strong persistence. For every country and year and unconditional on any fundamentals, the number of MEAs ratified of a specific type follows an autoregressive process which can be represented by an AR(1) process. In any case, the autoregressive process is as strong that we would not expect to be able to capture dynamics fully by the inclusion of explanatory variables. Accordingly, a country's MEA history should be used for explaining the current number of MEAs of that country, and it appears sufficient to include the first-order time lag of a country's number of MEAs for that purpose. Of course, lagged dependent variables are not exogenous in panel data models of limited time span. We follow Blundell et al. (2002) to



FIGURE 5 The Number of Multilateral Environmental Agreements in 2006 – Cluster: Seas

model the dynamics of the number of MEAs a country ratifies as a dynamic linear feedback model (LFM) for count data. In such a model, the conditional mean of a dependent count variable is assumed to be linear in the history of the process (see Chamberlain, 1993; Windmeijer and Santos Silva, 1997; Windmeijer, 2000, 2005, 2008; Hall, 2005).

The conditional mean in the standard LFM is defined as

$$E(y_{it}|y_{it-1}, x_{it}, \eta_i) = \gamma y_{it-1} + \exp(x'_{it}\beta + \eta_i)$$

= $\gamma y_{it-1} + \mu_{it}v_i,$ (1)

where y_{it} denotes the number of MEAs country *i*, *i* = 1, ..., *N*, has ratified in year *t*, *t* = 1, ..., *T*. x_{it} represents a vector of *K* explanatory variables, and $v_i \equiv \exp(\eta_i)$ is a scaling factor for the *i*-specific mean and $\mu_{it} = \exp(x'_{it}\beta)$. The parameters γ and β are to be estimated. According to Windmeijer (2008), the LFM can also be motivated as an entry–exit process with the probability of exit equal to $(1-\gamma)$. Moreover, the mean value for y_{it} is bounded from below by γy_{it-1} as $\mu_{it}v_i$ is non-negative.

As the number of MEAs is predetermined, that is, determined prior to the current period,

$$E(u_{it+j}y_{it}) = 0, \quad j \ge 0,$$

$$E(u_{it-s}y_{it}) \neq 0, \quad s \ge 1,$$

the within-group mean scaling estimator will be inconsistent in short panels. But with the Wooldridge's quasi-differencing transformation (Wooldridge, 1997),

$$q_{it} = \frac{y_{it} - \sum_{j=1}^{p} \gamma_j y_{it-j}}{\mu_{it}} - \frac{y_{it-1} - \sum_{j=1}^{p} \gamma_j y_{it-1-j}}{\mu_{it-1}}$$

the moment condition $E(q_{it}|y_{it-2}, x_{it-1}) = 0$ holds for predetermined variables. Furthermore, this transformation can handle endogenous x_{it} with $E(q_{it}|y_{it-2}, x_{it-2}) = 0$ as a valid moment condition. In either case, the generalised method of moments (GMM) estimator is defined as

$$\hat{\theta} = \arg\min\left(\frac{1}{N}\sum_{i=1}^{N}q_i(\theta)'Z_i\right)W_N^{-1}\left(\frac{1}{N}\sum_{i=1}^{N}Z_i'q_i(\theta)\right),$$

where $\theta = (\gamma_1, \dots, \gamma_p, \beta')'$, q_i is the T-p-1 vector of residuals q_{it} , Z_i is the matrix of instruments, and W_N is a weight matrix, which is optimally chosen as follows:

$$W_N(\hat{ heta}) = rac{1}{N} \sum_{i=1}^N Z_i' q_i(\hat{ heta}) q_i(\hat{ heta})' Z_i.$$

Here, $q_i(\hat{\theta})$ stems from an initial consistent estimate $\hat{\theta}$. Otherwise, for example, for the one-step GMM estimator, a proper initial weight matrix is

$$W_N = \frac{1}{N} \sum_{i=1}^N Z_i' Z_i.$$

By definition of $\hat{\theta}$, the GMM estimator is asymptotically normally distributed and its asymptotic variance can be computed as

$$\hat{var}(\hat{\theta}) = \frac{1}{N} (C(\hat{\theta})' W_N^{-1} C(\hat{\theta}))^{-1} C(\hat{\theta})' W_N^{-1} W_N(\hat{\theta}) W_N^{-1} C(\hat{\theta}) (C(\hat{\theta})' W_N^{-1} C(\hat{\theta}))^{-1},$$

where

$$C(\hat{\theta}) = \frac{1}{N} \sum_{i=1}^{N} \frac{\partial Z'_i q_i(\theta)}{\partial \theta} \Big|_{\hat{\theta}}.$$

In a second step, to capture spillover effects in the ratification across different MEA clusters, we extend the LFM using superscript c to refer to clusters.

$$E(y_{it}^{c}|y_{it-1}^{c}, x_{it}, \eta_{i}) = \gamma y_{it-1}^{c} + \exp(\tau y_{it-1}^{C \neq c} + x_{it}^{c}\beta + \eta_{i})$$

= $\gamma y_{it-1}^{c} + \mu_{it} v_{i},$ (2)

with $\exp(\tau y_{it-1}^{C\neq c} + x_{it}^{\prime}\beta)$ now. Here, the parameter τ measures the extent of spillovers of previously ratified MEAs in one or all other clusters (referred to by $C \neq c$) on the ratification of MEAs in cluster *c* conditional on the fundamentals in x_{it} .

We compare different strategies to estimate the parameters γ , τ and β . It turns out that, in the context of MEAs, the one-step estimator using Wooldridge's moment conditions is preferred over other GMM estimators. According to Windmeijer (2002), we found that the efficient two-step GMM estimator, which uses the estimates from the one-step estimator for the moments weighting matrix and the continuously updated GMM estimator that directly accounts for the dependence of the moments weighting matrix on the parameters in the optimisation (Hansen et al., 1996), is severely downward biased because of the relatively small number of countries *N* in our sample. Using a finite sample correction with block-bootstrapping in order to solve for the small sample bias could only reproduce the results from the one-step estimator (cf. Egger et al., 2011).

4. DETERMINANTS OF MEA RATIFICATION

Theoretical work on environmental agreements typically uses stylised emission abatement costs and benefits or climate change damage costs and side payments when determining partic-

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ipation in agreements. Other work applies a payoff share of the public good dubbed 'clean environment' (see Chandler and Tulkens, 1992; Hoel, 1992, 2005; Barrett, 1994, 1997, 2001; Hoel and Schneider, 1997; Tol, 1997, 2001; Carraro, 1998; Finus and Rundshagen, 1998a, 1998b; Frankel and Rose, 2002; Swanson and Mason, 2002; Barrett and Stavins, 2003; Carraro et al., 2003, 2006; Eyckmans and Tulkens, 2003; Lange and Vogt, 2003; Caparrós et al., 2004; Buchholz et al., 2005; Naghavi, 2005; Weikard et al., 2006). Since direct measures of the costs of pollution, of side payments and of payoff shares are not available, we follow the empirical literature to propose a reduced form of the determinants of environmental agreement participation as a function of observable economic, political and environmental fundamental variables. This specification is supposed to capture direct effects of these fundamentals on MEA ratification as well as indirect ones of the unmeasured costs of pollution, side payments and payoff shares. Table 1 lists the economic, political and environmental fundamental variables that are employed in our specification of the ratification of the number of MEAs. In what follows, we focus on a balanced panel data set where the dependent and the explanatory variables are available for 110 countries in all covered 47 years so that we are left with 5,170 observations.

In the horizontal dimension, the table provides the acronym of the respective variable, the average value, the standard deviation and the minimum and maximum of that variable. In the

Mean	SD	Min	Max
3			
32.2445	34.4753	0	212
4.0178	4.4381	0	27
3.8143	5.4397	0	30
2.8845	2.7611	0	21
8.5178	8.5266	0	48
12.9727	15.3181	0	94
28.2267	30.3859	0	185
28.4302	29.6939	0	186
29.3226	32.1034	0	195
23.7267	26.3782	0	164
19.2718	19.9618	0	123
23.5477	2.1554	17.8967	30.0656
9.3292	1.4516	6.2086	14.0895
1.4270	1.8332	-4.1154	5.2542
9.6680	18.5627	0	131
0.1818	0.3857	0	1
4.9579	1.9583	1	9.6
S			
52.8575	111.4035	0	683
3.6872	4.4801	-0.0197	27.7664
9.4035	13.1677	0	88.7417
	Mean 32.2445 4.0178 3.8143 2.8845 8.5178 12.9727 28.2267 28.4302 29.3226 23.7267 19.2718 23.5477 9.3292 1.4270 9.6680 0.1818 4.9579 52.8575 3.6872 9.4035	Mean SD S 32.2445 34.4753 4.0178 4.4381 3.8143 5.4397 2.8845 2.7611 8.5178 8.5266 12.9727 15.3181 28.2267 30.3859 28.4302 29.6939 29.3226 32.1034 23.7267 26.3782 19.2718 19.9618 23.5477 2.1554 9.3292 1.4516 1.4270 1.8332 9.6680 18.5627 0.1818 0.3857 4.9579 1.9583 15 52.8575 111.4035 3.6872 4.4801 9.4035 13.1677	Mean SD Min S 32.2445 34.4753 0 4.0178 4.4381 0 3.8143 5.4397 0 2.8845 2.7611 0 8.5178 8.5266 0 12.9727 15.3181 0 0 28.2267 30.3859 0 28.2267 30.3859 0 29.3226 32.1034 0 23.7267 26.3782 0 19.2718 19.9618 0 23.5477 2.1554 17.8967 9.3292 1.4516 6.2086 1.4270 1.8332 -4.1154 9.6680 18.5627 0 0.1818 0.3857 0 4.9579 1.9583 1 1.5 52.8575 111.4035 0 3.6872 4.4801 -0.0197 9.4035 13.1677 0 0 0.9435 0.9435

TABLE 1 Statistics of Balanced Data

Note:

(i) The sample is based on 110 countries, 47 years and 5,170 observations.

vertical dimension, we report statistics on the number of all MEAs at the top, on the number of MEAs by cluster in the second bloc at the top, on the number of MEAs in all but one cluster in the third bloc and on economic, political and environmental determinants in the fourth to sixth bloc, respectively.

The meaning of the acronyms in Table 1 is as follows. All variables under the heading 'Left-hand side (dependent) variables' measure numbers of MEAs across all or within a specific cluster as indicated. All variables under the heading 'Right-hand side spillover variables' measure numbers of MEAs in all clusters except the indicated one. The heading 'Economic determinants of MEAs' subsumes five covariates: LGDP measures a country's log real GDP in a year; LPOP measures a country's log population in a year; TRADE LIBERAL is a multilateral trade liberalisation measure for a country in a year in logs; INVEST LIBERAL is a multilateral investment liberalisation measure for a country in a year in logs; and LDC is a binary indicator variable which is unity for less developed countries and zero else. The only entry under the heading 'Political determinants of MEAs' is PFI that measures a country's political freedom index in a year. The heading 'Environmental determinants of MEAs' covers three covariates: PLANT SPECIES measures the number of endangered species in a country and year; CO₂ EMISSIONS measures a country's CO₂ emissions in kilo-tons (kt) in a year; and AGRRAW reflects agricultural raw materials exports in per cent of merchandise exports. The subsequent sections report some more detail on these variables and their hypothesised effects on the number of MEAs ratified per cluster.

a. Economic Determinants

The variables LGDP and LPOP are supposed to capture a country's economic size. The source data for these variables are taken from Maddison's (2003) historical time-series and extrapolated for more recent years by using indices of real GDP growth in US dollars and of population from the World Bank's World Development Indicators 2008 (World Bank, 2008), respectively.

TRADE LIBERAL measures the importance of bilateral and multilateral trade costs, among others, through preferential trade agreement (PTA) membership, to account for direct and indirect consequences of trade costs. In a nutshell, TRADE LIBERAL reflects the inverse of a country's multilateral trade costs in a year.¹ Similar to LGDP and LPOP, we use the log of inverted trade costs for TRADE LIBERAL. INVEST LIBERAL is a measure of a country's investment liberalisation that simply reflects the number of bilateral investment treaties (BITs) of a country.² Since TRADE LIBERAL or INVEST LIBERAL may be endogenous, we apply Wooldridge's quasi-differencing transformation – as stated above – which can deal with potentially endogenous regressors (see Wooldridge, 1997; Windmeijer, 2008).

Finally, LDC is included to allow for a link between the number of MEAs ratified and the development status of a country. This variable is constructed from data of the United Nations Office of the High Representative for the Least Developed Countries, Landlocked Developing Countries and Small Island Developing States (UN-OHRLLS).

¹ For more details on TRADE LIBERAL, see Anderson and van Wincoop (2003).

² We use BITs from the United Nations Conference of Trade and Development Treaty Database (UNCTAD, 2007).

	Number of MEAs	Biodiversity (Number of MEAs)	Atmosphere (Number of MEAs)	Land (Number of MEAs)	Chemicals (Number of MEAs)	Seas (Number of MEAs)
Lagged dependent var	riables					
y_{it-1}^{NBMEA}	0.1594***					
$y_{it-1}^{BIODIVERSITY}$		0.0058				
NBMEA≠BIODIVERSITY		0.0122***				
vATMOSPHERE		010122	0 3833***			
y_{it-1} NBMEA \neq ATMOSPHERE			0.0115***			
y_{it-1}			0.0115			
y_{it-1}^{LAND}				0.0264		
$y_{it-1}^{NBMEA \neq LAND}$				0.0020***		
$y_{it-1}^{CHEMICALS}$					0.0671***	
NBMEA≠CHEMICALS					0.0073***	
SEAS						0 0800***
y_{it-1} NBMEA \neq SEAS						0.0099
y_{it-1}	4.0					0.0099
L GDP.	0.8490***	0 5164***	0.4507***	0 1775***	0 6800***	0 5556***
	0.0901	0.0616	0.1053**	0.1775	-0.0397	0.1139
TRADE LIBERAL	0.6701***	0.4062***	0.5504***	0.1756*	0.5507***	0.3960***
INVEST LIBERAL _{it}	0.0032***	-0.0011	0.0008	0.0019**	0.0006	-0.0061***
LDC _{it}	0.8518*	0.7493**	0.0600	-0.0495	0.0169	-0.0410
Political determinants						
PFI _{it}	0.0620***	0.0173	0.0325*	-0.0013	-0.0409^{***}	0.0833***
Environmental determ	inants					
PLANT SPECIES _{it}	-0.0011^{***}	-0.0004^{**}	-0.0003	-0.0002	0.0002	-0.0014^{***}
AGRRAW _{it}	-0.0180***	-0.0088**	-0.0079	0.0017	-0.0041	-0.0185***
$CO_2 EMISSIONS_{it}$	-0.0117	-0.0163	0.0495***	-0.0082^{*}	-0.0298	0.0044

TABLE 2
Parameter Estimates with $E(y_{it}^c y_{it-1}^c, x_{it}, v_i)$

Notes:

(i) *t*-Statistics in parentheses.

(ii) There are 110 countries and 5,170 observations in all six ONE-STEP regressions.

(iii) The parameters are estimated over the period 1962-2006.

(iv) Once and twice lagged levels of the dependent and the independent variables are used as instruments (i.e. values of 1960 and 61 are used as instruments for 1962).

(v) *, **, *** indicates that parameters are significant at 10%, 5% and 1%, respectively.

b. Political Determinants

The index of political freedom (PFI) is published by the Fraser Institute (Gwartney et al., 2007) and ranges from 1 to 10. The index is a compound measure of the quality of a country's legal structure, the security of property rights and the ease of access to sound money in a year. A higher index value reflects greater political freedom in a country and year.

c. Environmental Determinants

The three environmental determinants of MEA ratification are based on the following sources. PLANT SPECIES reflects the degree of habitat protection of species (the number of

	Biodiversity (Number of MEAs)				
Lagged dependent varia	ables				
VBIODIVERSITY	0.0058	0.0138	0.0139	0.0100	0.0129
$y_{it-1}^{NBMEA \neq BIODIVERSITY}$ $y_{it-1}^{NBMEA \neq BIODIVERSITY}$	0.0122***				
$y_{it-1}^{ATMOSPHERE}$		0.0125**			
y_{it-1}^{LAND}			0.0355*		
$y_{it-1}^{CHEMICALS}$				0.0316***	
y_{it-1}^{SEAS}					0.0185***
Economic determinants					
LGDP _{it}	0.5164***	0.7391***	0.6095***	0.5637***	0.5471***
LPOP _{it}	0.0616	0.0414	0.0727	0.0630	0.0544
TRADE LIBERAL _{it}	0.4062***	0.5231***	0.5211***	0.3952***	0.4459***
INVEST LIBERAL _{it}	-0.0011	0.0029**	0.0048***	0.0008	0.0030**
LDC _{it}	0.7493**	0.9797**	0.9099**	0.8675**	0.7793**
Political determinants					
PFI _{it}	0.0173	0.0325***	0.0297**	0.0363***	0.0125
Environmental determin	nants				
PLANT SPECIES _{it}	-0.0004^{**}	-0.0004^{**}	-0.0004^{**}	-0.0006^{***}	-0.0002
AGRRAW,	-0.0088**	-0.0110^{***}	-0.0118***	-0.0099***	-0.0081**
CO ₂ EMISSIONS _{it}	-0.0163	-0.0178	-0.0150	-0.0163	-0.0147

	TABLE 3	
Parameter Estimates	with $E(y_{it}^c y_{it-1}^c, x_{it}, v_i); c =$	BIODIVERSITY

(i) t-Statistics in parentheses.

(ii) There are 110 countries and 5,170 observations in all six ONE-STEP regressions.

(iii) The parameters are estimated over the period 1962-2006.

(iv) Once and twice lagged levels of the dependent and the independent variables are used as instruments (i.e. values of 1960 and 61 are used as instruments for 1962).

(v) ** and *** indicates that parameters are significant at 5% and 1%, respectively.

endangered species in a country and year) as a key measure of *Biodiversity* in a country (Daniel et al., 2008). The variable is published in the World Bank's World Development Indicators 2008 (World Bank, 2008) and linearly interpolated for missing years. The variable takes on the highest value (of 683) for Malaysia in 2004. The number of endangered species is <5 for 51 per cent of the countries and years. CO_2 EMISSIONS – measured in kilo-tons – also comes from the World Bank's World Development Indicators 2008 (World Bank, 2008).³ There is one country-year observation for which CO_2 EMISSIONS takes on a negative value, Senegal in 1968. Otherwise, CO_2 EMISSIONS is positive throughout. In the sample covered, it takes on the highest value for China in 2006.

AGRRAW – reflecting agricultural raw materials exports in per cent of merchandise exports – is based on the World Bank's World Development Indicators 2008 (World Bank, 2008) and ranges from 0 to 88.74 per cent, with an average of 9.4 per cent. In the covered

³ We ran alternative regressions using CO_2 emissions *per capita* and CO_2 emissions per GDP, and the corresponding results were very similar to the ones based on the proposed measure.

	Atmosphere (Number of MEAs)				
Lagged dependent varia	bles				
VATMOSPHERE	0.3833***	0.4944***	0.5077***	0.4248***	0.3714***
$y_{it-1}^{NBMEA \neq ATMOSPHERE}$ $y_{it-1}^{NBMEA \neq ATMOSPHERE}$	0.0115***				
$y_{it-1}^{BIODIVERSITY}$		0.0180			
y_{it-1}^{LAND}			0.0168		
$y_{it-1}^{CHEMICALS}$				0.0295***	
y_{it-1}^{SEAS}					0.0266***
Economic determinants					
LGDP _{it}	0.4507***	0.5343***	0.4232***	0.4218***	0.4475***
LPOP _{it}	0.1053**	0.0783	0.0808	0.1252**	0.0934*
TRADE LIBERAL _{it}	0.5504***	0.5491***	0.4710***	0.5045***	0.5582***
INVEST LIBERAL _{itit}	0.0008	0.0044**	0.0061***	0.0021	0.0011
LDC	0.0600	0.2053	0.1996	0.1238	0.0766
Political determinants					
PFI _{it}	0.0325*	0.0371	0.0523**	0.0616***	0.0310*
Environmental determine	ants				
PLANT SPECIES _{it}	-0.0003	-0.0004	-0.0007^{**}	-0.0006^{*}	-0.0002
AGRRAW _{it}	-0.0079	-0.0089	-0.0112	-0.0084	-0.0075
CO ₂ EMISSIONS _{it}	0.0495**	0.0589**	0.0582**	0.0613**	0.0399

Parameter Estimates with $E(y_{it}^c|y_{it-1}^c, x_{it}, v_i)$; c=ATMOSPHERE

Notes:

(i) t-Statistics in parentheses.

(ii) There are 110 countries and 5,170 observations in all six ONE-STEP regressions.

(iii) The parameters are estimated over the period 1962-2006.

(iv) Once and twice lagged levels of the dependent and the independent variables are used as instruments (i.e. values of 1960 and 61 are used as instruments for 1962).

(v) *, **, *** indicates that parameters are significant at 10%, 5% and 1%, respectively.

data, Nepal and Singapore display the highest values of AGRRAW in 1960, while Burkina Faso, Benin and the Central African Republic display the highest value in 2006.⁴

5. RESULTS

We organise the presentation of results in two subsections. While Section 5a summarises parameter estimates and their standard errors, Section 5b is devoted to the discussion of associated marginal effects.

⁴ We employed the land area (in per cent of total land area), the forest area of a country (in per cent of total land area or in square kilometres) and a country's permanent cropland (in per cent of total land area) from the World Bank's World Development Indicators 2008 (World Bank, 2008) as possible alternatives to AGRRAW. However, these variables are much more collinear with the included covariates (especially, with CO_2 EMISSIONS and PLANT SPECIES) than AGRRAW is. Therefore, we prefer using AGRRAW instead of those alternatives.

	Land (Number of MEAs)				
Lagged dependent var	riables				
VLAND	0.0264	0.0269	0.0263	0.0288	0.0327
$y_{it-1}^{NBMEA \neq LAND}$ $y_{it-1}^{NBMEA \neq LAND}$	0.0020***				
$y_{it-1}^{BIODIVERSITY}$		0.0068***			
$y_{it-1}^{ATMOSPHERE}$			0.0092***		
$y_{it-1}^{CHEMICALS}$				0.0046**	
y_{it-1}^{SEAS}					0.0016*
Economic determinan	ts				
LGDP _{it}	0.1775***	0.2155***	0.2230***	0.1988***	0.2282***
LPOP _{it}	0.2521**	0.2321**	0.2457***	0.2546**	0.2256**
TRADE LIBERAL _{it}	0.1756*	0.1868^{*}	0.1761*	0.1778^{*}	0.1971*
INVEST LIBERAL _{it}	0.0019**	0.0027***	0.0012	0.0023**	0.0030***
LDC _{it}	-0.0495	-0.0119	0.0430	-0.0094	0.0023
Political determinants	1				
PFI _{it}	-0.0013	0.0003	0.0014	0.0033	0.0013
Environmental determ	vinants				
PLANT SPECIES _{it}	-0.0002	-0.0003	-0.0003	-0.0003^{*}	-0.0003
AGRRAW _{it}	0.0017	0.0006	0.0006	0.0008	0.0012
CO ₂ EMISSIONS _{it}	-0.0082^{*}	-0.0088^{*}	-0.0088^{**}	-0.0088^{**}	-0.0106^{**}

	TAB	SLE 5
Parameter Estimates	with	$E(y_{it}^c y_{it-1}^c, x_{it}, v_i); c=LAND$

(i) *t*-Statistics in parentheses.

(ii) There are 110 countries and 5,170 observations in all six ONE-STEP regressions.

(iii) The parameters are estimated over the period 1962-2006.

(iv) Once and twice lagged levels of the dependent and the independent variables are used as instruments (i.e. values of 1960 and 61 are used as instruments for 1962).

(v) *, **, *** indicates that parameters are significant at 10%, 5% and 1%, respectively.

a. Parameter Estimates

The models in Table 2 include the economic, political and environmental determinants along with the lagged dependent number of MEAs ratified. There are six columns: one for all MEAs and one for each of the five clusters. The associated results can be summarised as follows. First, the direction of a specific determinant's effect on the number of MEAs tends to be the same (with the occasional exception): higher levels of the lagged dependent variable, economic determinants and political (benefit) determinants affect the number of MEAs ratified positively, whereas environmental (cost) determinants, if they are significant, exert a negative impact on the number of MEAs. Second, the coefficients of the lagged dependent variable are positive and significantly different from zero throughout. Accordingly, adjustment dynamics are important for MEA ratification within and across clusters.

Third, LGDP and TRADE LIBERAL exert a large, positive and significant impact for all MEAs and all MEA clusters. Hence, economic size and multilateral trade liberalisation are the most important and robust drivers of MEA ratification within and across clusters. Fourth, the impact of investment liberalisation (INVEST LIBERAL), of political freedom (PFI) and

	Chemicals	Chemicals	Chemicals	Chemicals	Chemicals
	(Number of MEAs)				
Lagged dependent varia	ables				
$y_{it=1}^{CHEMICALS}$	0.0671***	0.0739***	0.0723***	0.0774***	0.0648***
$y_{it-1}^{NBMEA \neq CHEMICALS}$	0.0073***				
$y_{it-1}^{BIODIVERSITY}$		0.0060			
$y_{it-1}^{ATMOSPHERE}$			0.0254***		
y_{it-1}^{LAND}				0.0059	
V ^{SEAS}					0.0115***
<i>Economic determinants</i>					
LGDP _{it}	0.6899***	0.7467***	0.7137***	0.7229***	0.7032***
LPOP _{it}	-0.0397	-0.0287	-0.0275	-0.0292	-0.0382
TRADE LIBERAL _{it}	0.5507***	0.5827***	0.5443***	0.5705***	0.5735***
INVEST LIBERAL _{it}	0.0006	0.0048***	0.0002	0.0050***	0.0023**
LDC _{it}	0.0169	0.0841	0.0987	0.0756	0.0586
Political determinants					
PFI _{it}	-0.0409^{***}	-0.0245^{*}	-0.0278^{**}	-0.0175	-0.035^{***}
Environmental determin	ants				
PLANT SPECIES _{it}	0.0002	0.0001	0.0001	0.0000	0.0002
AGRRAW _{it}	-0.0041	-0.0050	-0.0053	-0.0053	-0.0042
CO ₂ EMISSIONS _{it}	-0.0298**	-0.0261^{*}	-0.0261^{*}	-0.0270^{**}	-0.0297^{**}

TABLE 6 Parameter Estimates with $E(y_{it}^c|y_{it-1}^c, x_{it}, v_i); c=CHEMICALS$

Notes:

(i) t-Statistics in parentheses.

(ii) There are 110 countries and 5,170 observations in all six ONE-STEP regressions.

(iii) The parameters are estimated over the period 1962-2006.

(iv) Once and twice lagged levels of the dependent and the independent variables are used as instruments (i.e. values of 1960 and 61 are used as instruments for 1962).

(v) *, **, *** indicates that parameters are significant at 10%, 5% and 1%, respectively.

of emissions (CO₂ EMISSIONS) is often significantly different from zero but differs qualitatively across clusters.

Notice that the specifications in Table 2 do not distinguish between spillover effects associated with the ratification of MEAs from different clusters. In a next step, we shed light on such spillover effects by including the lagged dependent variable as in Table 2 along with the respective spillover variable as introduced in the second vertical bloc of Table 1. The respective results for each cluster are summarised in Tables 3–7. Hence, in addition to the impact of all MEAs ratified in clusters other than c on the number of MEAs in c, we also shed light on the role of spillovers from specific non-c clusters in Tables 3–7.

Tables 3–7 suggest the following conclusions. First, dynamic adjustment within a cluster is found to be important for all clusters except biodiversity and land. Spillover effects from all other clusters are found to be important throughout. Hence, there is contagion in the ratification of MEAs across issues. Economic size (LGDP) and multilateral trade liberalisation (TRADE LIBERAL) are found to be key drivers of the ratification of MEAs across the board also when conditioning on cross-cluster spillovers. Investment liberalisation appears to be

	Seas (Number of MEAs)				
Lagged dependent var	iables				
y_{it-1}^{SEAS}	0.0899***	0.0903***	0.0947***	0.0924***	0.0914***
$NBMEA \neq SEAS$ Y_{it} 1	0.0099***				
$y_{it-1}^{BIODIVERSITY}$		0.0303***			
$y_{it=1}^{ATMOSPHERE}$			0.0072		
y_{it-1}^{LAND}				0.0121	
y _{it-1}					0.0161**
Economic determinant	's				
LGDP _{it}	0.5556***	0.5790***	0.6448***	0.6439***	0.5779***
LPOP _{it}	0.1139	0.1283	0.1261	0.1367	0.1245
TRADE LIBERAL _{it}	0.3960***	0.4274***	0.4634***	0.4896***	0.4099***
INVEST LIBERAL _{it}	-0.0061^{***}	-0.0033^{*}	-0.0023^{*}	-0.0019	-0.0047^{***}
LDC _{it}	-0.0410	-0.0464	0.0501	0.0491	0.0074
Political determinants					
PFI _{it}	0.0833***	0.0791***	0.0842***	0.0841***	0.0840***
Environmental determ	inants				
PLANT SPECIES _{it}	-0.0014^{***}	-0.0014^{***}	-0.0012^{***}	-0.0012^{***}	-0.0014^{***}
AGRRAWit	-0.0185^{***}	-0.0189^{***}	-0.0177^{***}	-0.0181^{***}	-0.0182^{***}
CO ₂ EMISSIONS _{it}	0.0044	0.0061	0.0088	0.0036	0.0057

TABLE 7 Parameter Estimates with $E(y_{it}^c|y_{it-1}^c, x_{it}, v_i); c=SEAS$

(i) t-Statistics in parentheses.

(ii) There are 110 countries and 5,170 observations in all six ONE-STEP regressions.

(iii) The parameters are estimated over the period 1962-2006.

(iv) Once and twice lagged levels of the dependent and the independent variables are used as instruments (i.e. values of 1960 and 61 are used as instruments for 1962).

(v)* and *** indicates that parameters are significant at 10%, and 1%, respectively.

important for the clusters land and seas: while it stimulates ratifying MEAs in the former, it appears to deter ratifying MEAs in the latter.

Among the environmental fundamentals, a larger number of endangered species deters MEA ratification in particular for the clusters biodiversity and seas. A bigger percentage of agricultural raw materials exports in merchandise exports tends to reduce the incentive to conclude MEAs in the clusters biodiversity and seas. Interestingly, CO₂ EMISSIONS deters MEAs in the clusters chemicals and hazardous wastes and land, while they appear to stimulate MEAs in the cluster atmosphere.

b. Marginal Effects

The parameter estimates of the regression results in Tables 2–7 offer insights into qualitative, but not the quantitative, effects of the fundamental variables on MEA ratification. Let us augment this information by reporting a set of marginal effects in Table 8. In the interest of space, let us report marginal effects only for statistically significant parameters in the last five columns of Table 2 (corresponding also to the first columns of Tables 3–7). Moreover, let us report effects to one-standard deviation increases in the respective fundamental variables. In

	Biodiversity	Atmosphere	Land	Chemicals	Seas
Lagged dependent varia	bles				
NBMEA≠BIODIVERSITY	44 88				
y_{it-1} NBMEA \neq ATMOSPHERE	00	10.70			
y_{it-1}		40.70			
$y_{it-1}^{NBMEA \neq LAND}$			6.63		
NBMEA≠CHEMICALS				21.24	
J_{it-1} NBMEA \neq SEAS				21.21	21.05
y_{it-1}					21.85
Economic determinants					
LGDP _{it}	204.36	164.17	46.61	342.39	231.20
LPOP _{it}	n/v	16.52	44.19	n/v	n/v
TRADE LIBERAL _{it}	110.57	174.28	37.98	174.43	106.67
INVEST LIBERAL _{it}	n/v	n/v	3.59	n/v	-10.71
LDC _{it}	33.51	n/v	n/v	n/v	n/v
Political determinants					
PFI _{it}	n/v	6.57	n/v	-7.53	17.72
Environmental determine	ants				
PLANT SPECIES _{it}	-4.36	n/v	n/v	n/v	-14.44
AGRRAW _{it}	-10.94	n/v	n/v	n/v	-21.62
CO ₂ EMISSIONS _{it}	n/v	24.83	-3.61	-12.50	n/v

TABLE 8 Marginal Effects of Explanatory Variables

(i) Figures are percentage changes of MEA counts in response to a one-standard deviation increase in an explanatory variable. Marginal effects are calculated only for statistically significant parameters in the last five columns of Table 2. n/v means *no value* and fills the cells where the corresponding β -coefficients are not significant.

Table 8, we use n/v to indicate 'no value' for statistically insignificant coefficients in the last five columns of Table 2.

The marginal effects may be summarised as follows. First, economic size (LGDP) and multilateral trade liberalisation (TRADE LIBERAL) are the most important drivers of MEA ratification across the board. For instance, a one-standard deviation change in LGDP raises the number of MEAs ratified by in between about 47 (land) and 342 (chemicals and hazard-ous wastes). A one-standard deviation change in TRADE LIBERAL exerts a somewhat smaller effect on the number of MEAs ratified in all clusters except atmosphere. In comparison, the remaining marginal effects are negligible.

6. CONCLUSION

This paper analyses the impact of economic, political and environmental fundamentals on the ratification of multilateral environmental agreements (MEAs) in five different clusters of issues: biodiversity, atmosphere, land, chemicals and hazardous wastes, and seas. The results point to an overwhelming importance of economic size and multilateral trade liberalisation as drivers of MEA ratification across clusters. Other determinants have much smaller and in part even qualitatively ambiguous effects on the number of MEAs concluded. Apart from the mentioned fundamentals, adjustment costs and, even more so, spillovers from the ratification of MEAs across clusters are found to be important for the scope of participation in MEAs.

REFERENCES

- Alpay, E., S. Buccola and J. Kerkvliet (2002), 'Productivity Growth and Environmental Regulation in Mexican and US Food Manufacturing', *American Journal of Agricultural Economics*, 84, 4, 887–901.
- Anderson, J. E. and E. van Wincoop (2003), 'Gravity with Gravitas: A Solution to the Border Puzzle', *American Economic Review*, **93**, 1, 170–92.
- Barrett, S. (1994), 'Self-enforcing International Environmental Agreements', Oxford Economic Papers, 46, 1, 874–94.
- Barrett, S. (1997), 'Heterogeneous International Environmental Agreements', in C. Carraro (ed.), International Environmental Agreements: Strategic Policy Issues (Cheltenham: Edward Elgar), 9–25.
- Barrett, S. (2001), 'International Cooperation for Sale', European Economic Review, 45, 10, 1835–50.
- Barrett, S. and R. Stavins (2003), 'Increasing Participation and Compliance in International Climate Change Agreements', International Environmental Agreements: Politics, Law and Economics, 3, 4, 349–76.
- Berman, E. and L. T. M. Bui (2001), 'Environmental Regulation and Productivity: Evidence From Oil Refineries', *Review of Economics and Statistics*, 83, 3, 498–510.
- Beron K. J., J. C. Murdoch and W. P. M. Vijverberg (2003), 'Why Cooperate? Public Goods, Economic Power, and The Montreal Protocol', *Review of Economics and Statistics*, **85**, 2, 286–97.
- Blundell R., R. Griffith and F. A. G. Windmeijer (2002), 'Individual Effects and Dynamics in Count Data Models', *Journal of Econometrics*, 108, 1, 113–31.
- Buchholz, W., A. Hauptt and W. Peters (2005), 'International Environmental Agreements and Strategic Voting', Scandinavian Journal of Economics, 107, 1, 175–95.
- Bjørn, E., R. Golombek and A. Raknerud (1997), 'Environmental Regulation and Plant Exit: A Logit Analysis Based on Establishment Panel Data', *Environmental and Resource Economics*, 11, 1, 35–59.
- Caparrós A., A. Hammoudi and T. Tazdaït (2004), 'On Coalition Formation with Heterogeneous Agents', *Fondazione Eni Enrico Mattei Nota di Lavoro Series* 70.2004 (Rome: Fondazione Eni Enrico Mattei).
- Carraro C. (1998), 'Beyond Kyoto. A Game-Theoretic Perspective', paper prepared for the OECD workshop on *Climate Change and Economic Modelling. Background Analysis for the Kyoto Protocol*, Paris, 17–18 October 1998.
- Carraro C., J. Eyckmans and M. Finus (2006), 'Optimal Transfers and Participation Decisions in International Environmental Agreements', *Review of International Organisations*, 1, 4, 379–96.
- Carraro C., C. Marchiori and S. Oreffice (2003), 'Endogenous Minimum Participation in International Environmental Treaties', *Fondazione Eni Enrico Mattei Nota di Lavoro Series* 113.2003 (Rome: Fondazione Eni Enrico Mattei).
- Center for International Earth Science Information Network (CIESIN) (2006), Socioeconomic Data and Applications Center (SEDAC) *Environmental Treaties and Resource Indicators*. Available at: Http://sedac.ciesin.columbia.edu/entri (accessed 1 August 2010).
- Chamberlain, G. (1993), Feedback in Panel Data Models, Discussion Paper No. 1656 (Cambridge, MA: Harvard Institute of Economic Research).
- Chandler, P. and H. Tulkens (1992), 'Theoretical Foundations of Negotiations and Cost Sharing in Transfrontier Pollution Problems', *European Economic Review*, **36**, 2–3, 388–98.
- Daniel, C. E., M. A. Levy, C. H. Kim, A. de Sherbinin, T. Srebotnjak and V. Mara (2008), Environmental Performance Index (New Haven, CT: Yale Center for Environmental Law and Policy).
- d'Arge, R. (1974), 'International Trade, Domestic Income, and Environmental Controls: Some Empirical Estimates', in A. Kneese (ed.), Managing the Environment: International Economic Cooperation for Pollution Control (New York: Praeger), 289–315.
- Egger, P. H., C. Jessberger and M. Larch (2011), 'Trade and Investment Liberalisation as Determinants of Multilateral Environmental Agreement Membership', *International Tax and Public Finance*, **18**, 6, 605–33.
- Eyckmans J. and H. Tulkens (2003), 'Simulating Coalitionally Stable Burden Sharing Agreements for the Climate Change Problem', *Resource and Energy Economics*, **25**, 4, 299–327.
- Finus, M. and B. Rundshagen (1998a) 'Renegotiation-proof Equilibria in a Global Emission Game When Players Are Impatient', *Environmental and Resource Economics*, **12**, 3, 275–306.
- Finus M. and B. Rundshagen (1998b), 'Toward a Positive Theory of Coalition Formation and Endogenous Instrumental Choice in Global Pollution Control', *Public Choice*, 96, 1–2, 145–86.

- Frankel, J. A. and A. K. Rose (2002), 'Is Trade Good or Bad for the Environment? Sorting Out the Causality', Working papers 9201 (Cambridge, MA: National Bureau of Economic Research).
- Gwartney, J., R. Lawson, R. S. Sobel and P. T. Leeson (2007), Economic Freedom of the World: 2007 Annual Report (Vancouver: The Fraser Institute). Available at: Http://www.freetheworld.com/ (accessed 1 August 2010).
- Hall, A. R. (2005), Generalised Method of Moments (Oxford: Oxford University Press).
- Hansen, L. P., J. Heaton and A. Yaron (1996), 'Finite Sample Properties of Some Alternative GMM Estimators Obtained from Financial Market Data', *Journal of Business and Economic Statistics*, 14, 3, 262–80.
- Hoel, M. (1992), 'International Environmental Conventions: The Case of Uniform Reductions of Emissions', Environmental and Resource Economics, 2, 2, 141–59.
- Hoel, M. (2005), 'The Triple Inefficiency of Uncoordinated Environmental Policies', Scandinavian Journal of Economics, 107, 1, 157–73.
- Hoel, M. and K. Schneider (1997), 'Incentives to Participate in an International Environmental Agreement', *Environmental and Resource Economics*, **9**, 2, 153–70.
- Lange, A. and C. Vogt (2003) 'Cooperation in International Environmental Negotiations due to a Preference for Equity', *Journal of Public Economics*, 87, 9–10, 2049–67.
- Lanoie, P., M. Patry and R. Lajeunesse (2008), 'Environmental Regulation and Productivity: Testing the Porter Hypothesis', *Journal of Productivity Analysis*, **30**, 2, 121–28.
- Leonard, H. J. (1988), Pollution and the Struggle for World Product (Cambridge: Cambridge University Press).
- Maddison, A. (2003), The World Economy: Historical Statistics (Paris: OECD).
- Markusen, J. R. (1999), 'Location Choice, Environmental Quality and Public Policy', in J. C. J. M. van den Bergh (ed.), *Handbook of Environmental and Resource Economics* (Cheltenham: Edward Elgar), 569–80.
- McGuire, M. (1982), 'Regulation, Factor Rewards, and International Trade', *Journal of Public Economics*, **17**, 3, 335–54.
- Millimet, D. L. and J. A. List (2004), 'The Case of the Missing Pollution Haven Hypothesis', *Journal of Regulatory Economics*, **26**, 3, 239–62.
- Mitchell, R. (2003), 'International Environmental Agreements: A Survey of Their Features, Formation and Effects', Annual Review of Environment and Resources, 28, 429–61.
- Mitchell, R. (2007), 'International Environmental Agreements Database Project', version 2007.1, 2002–2007. Available at: Http://iea.uoregon.edu (accessed 1 August 2010).
- Murdoch J. C., T. Sandler and P. M. Vijverberg (2003) 'The Participation Decision Versus the Level of Participation in an Environmental Treaty: A Spatial Probit Analysis', *Journal of Public Economics*, 87, 2, 337–62.
- Naghavi, E. (2005), 'Multilateral Environmental Agreements and Trade Obligations: A Theoretical Analysis of the Doha Proposal', *Fondazione Eni Enrico Mattei Nota di Lavoro Series* 52.2005 (Rome: Fondazione Eni Enrico Mattei).
- Neumayer, E. (2002), 'Do Democracies Exhibit Stronger International Environmental Commitment? A Cross-country Analysis', *Journal of Peace Research*, **39**, 2, 139–64.
- OECD (1985), *The Macro-Economic Impact of Environmental Expenditures* (Paris: Organisation for Economic Cooperation and Development).
- Pethic, R. (1976), 'Pollution, Welfare, and Environmental Policy in the Theory of Comparative Advantage', *Journal of Environmental Economics and Management*, **2**, 3, 160–69.
- Pearson, C. (1985), Down to Business: Multinational Corporations, the Environment and Development (Washington, DC: World Resources Institute).
- Pearson, C. (1987), Multinational Corporation, the Environment and Development (Washington, DC: World Resources Institute).
- Porter, M. E. (1991), 'America's Green Strategy', Scientific American, 264, 4, 96.
- Porter, M. E. and C. van der Linde (1995) 'Toward a New Conception of the Environment-competitiveness Relationship', *Journal of Economic Perspectives*, **9**, 4, 97–118.
- Rose A. K. and M. M. Spiegel (2009), 'Non-economic Engagement and International Exchange: The Case of Environmental Treaties', *Journal of Money, Credit, and Banking*, **41**, 2–3, 337–63.
- Siebert, H. (1977), 'Environmental Quality and the Gains from Trade', Kyklos, 30, 4, 651–73.

- Swanson, T. and R. Mason (2002), 'The Impact of International Environmental Agreements: The Case of the Montreal Protocol', *Fondazione Eni Enrico Mattei Nota di Lavoro Series* 81.2002 (Rome: Fondazione Eni Enrico Mattei).
- Taylor, M. S. (2005), 'Unbundling the Pollution Haven Hypothesis', Advances in Economic Analysis and Policy, 4, 2, 1408–34.
- Tobey, J. A. (1990), 'The Effects of Domestic Environmental Policies on Patterns of World Trade: An Empirical Test', *Kyklos*, **43**, 2, 191–209.
- Tol, R. S. J. (1997), 'On the Optimal Control of Carbon Dioxide Emissions: An Application of FUND', Environmental Modeling and Assessment, 2, 3, 151–63.
- Tol, R. S. J. (2001), 'Climate Coalitions in an Integrated Assessment Model', *Computational Economics*, **18**, 2, 159–72.
- Ulph, A. (1999), 'Strategic Environmental Policy and Foreign Trade', in J. C. J. M. van den Bergh (ed.), Handbook of Environmental and Resource Economics (Cheltenham: Edward Elgar), 433–48.
- UNCTAD (2007), World Investment Report Transnational Corporations, Extractive Industries and Development (New York and Geneva: United Nations Conference on Trade and Development).
- UNEP (2001), 'Multilateral Environmental Agreements: A Summary', background paper presented by the secretariat, *United Nations Environment Programme*, UNEP /IGM /1 /INF /1, 30 March 2001. Available at: Http://www.unep.org/IEG/docs/working (accessed 1 August 2010).
- United Nations Office of the High Representative for the Least Developed Countries, Landlocked Developing Countries and Small Island Developing States (UN-OHRLLS). Available at: Http://www.u-n.org/special-rep/ohrlls/ldc/list.htm (accessed 1 August 2010).
- van Beers, C. and J. C. J. M. van den Bergh (1997), 'An Empirical Multi-Country Analysis of the Impact of Environmental Regulations on Foreign Trade Flows', *Kyklos*, **50**, 1, 29–46.
- van Beers, C. and J. C. J. M. van den Bergh (2000), The Impact of Environmental Policy on Foreign Trade, Discussion Paper 069, 1–15 (Amsterdam: Tinbergen Institute).
- Walter, I. (1982), 'Environmentally Induced Industrial Relocation to Developing Countries', in S. Rubin and T. Graham (eds.), *Environment and Trade* (Montclair, NJ: Allenheld Osman and Co.), 67–101.
- Weikard, H.-P., M. Finus and J.-C. Altamirano-Cabrera (2006), 'The Impact of Surplus Sharing on the Stability of International Climate Agreements', Oxford Economic Papers, 58, 2, 209–32.
- Windmeijer, F. A. G. (2000), 'Moment Conditions for Fixed Effects Count Data Models with Endogenous Regressors', *Economics Letters*, 68, 1, 21–24.
- Windmeijer, F. A. G. (2002), 'ExpEnd, A Gauss Programme for Non-linear GMM Estimation of *EXP*onential Models with *END*ogenous Regressors for Cross Section and Panel Data', CEMMAP Working Paper CWP 14/02 (London: Institute for Fiscal Studies).
- Windmeijer, F. A. G. (2005), 'A Finite Sample Correction for the Variance of Linear Efficient Two-step GMM Estimators', *Journal of Econometrics*, **126**, 1, 25–51.
- Windmeijer, F. A. G. (2008), 'GMM for Panel Count Data Models', in M. Laszlo and P. Sevestre (eds.), The Econometrics of Panel Data, 3rd edn (New York: Springer), 603–24.
- Windmeijer, F. A. G. and J. M. Santos Silva (1997), 'Endogeneity in Count Data Models: An Application to Demand for Health Care', *Journal of Applied Econometrics*, 12, 3, 281–94.
- Wooldridge, J. M. (1997), 'Multiplicative Panel Data Models without the Strict Exogeneity Assumption', *Econometric Theory*, 13, 5, 667–78.
- World Bank (2008), World Development Indicators (Washington, DC: World Bank Group).
- World Resources Institute (2003), 'Chapter 7. International Environmental Governance', in World Resources Institute (ed.), World Resources 2002–2004: Decisions for the Earth: Balance, Voice, and Power (Washington, DC: World Resource Institute), 137–72.
- Xu, X. and L. Song (2000), 'Regional Cooperation and the Environment: Do 'Dirty' Industries Migrate?', *Review of World Economics (Weltwirtschaftliches Archiv)*, **136**, 1, 137–57.